

# Kinship Structure and Maternal and Child Mortality in Croatia 1750-1898

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# Study Area

- Seven parishes in central Croatia (Slavonia)
- Retaken from the Ottomans in 1683, and repopulated by the Hapsburgs beginning in 1700
- Civil/Military peasantry - Croats (and Serbs) were the prime cannon fodder of the Austro-Hungarian Empire
- Subsistence agriculture and slow capitalist development
- “New Feudalism”

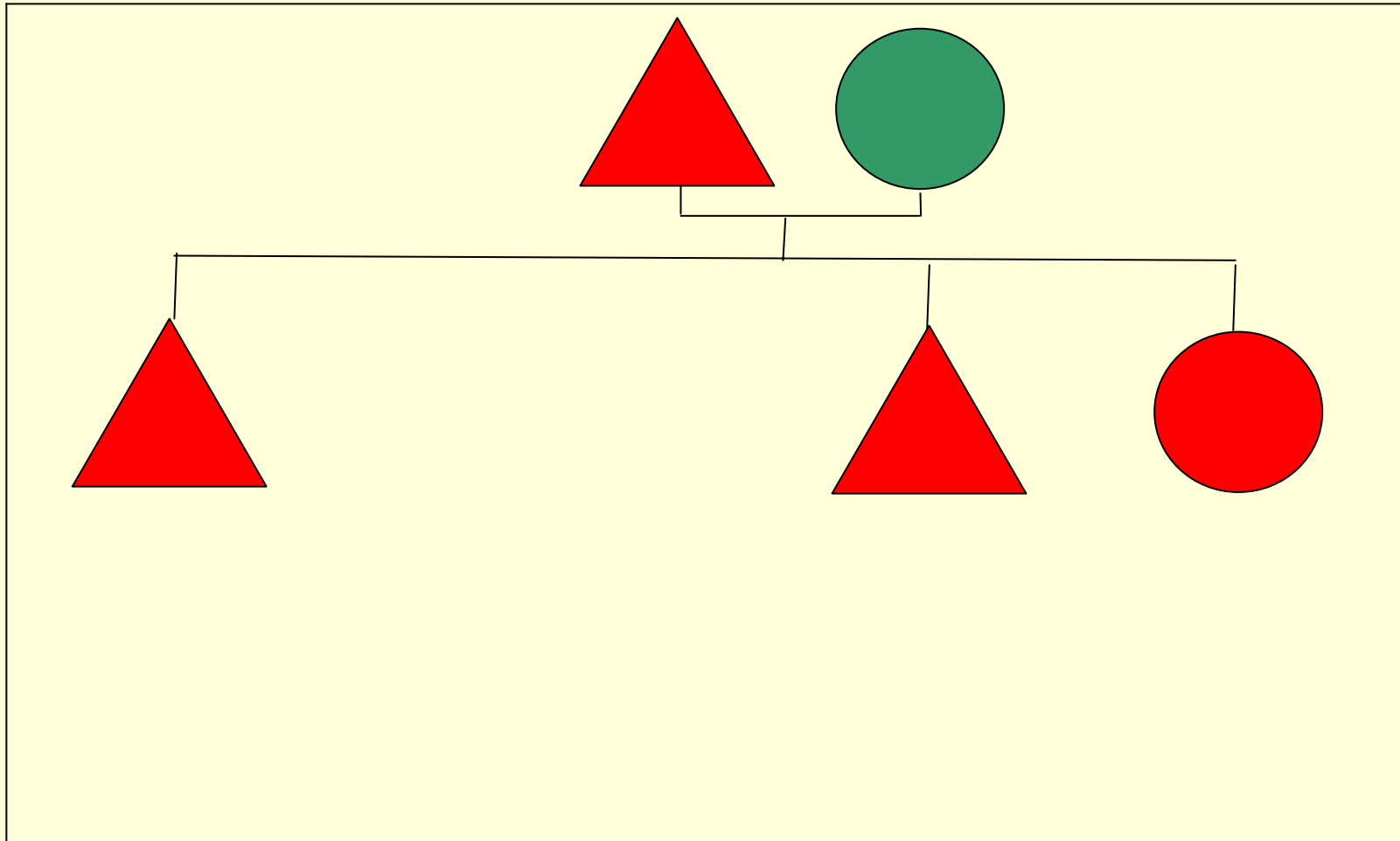
Fig. 11.1



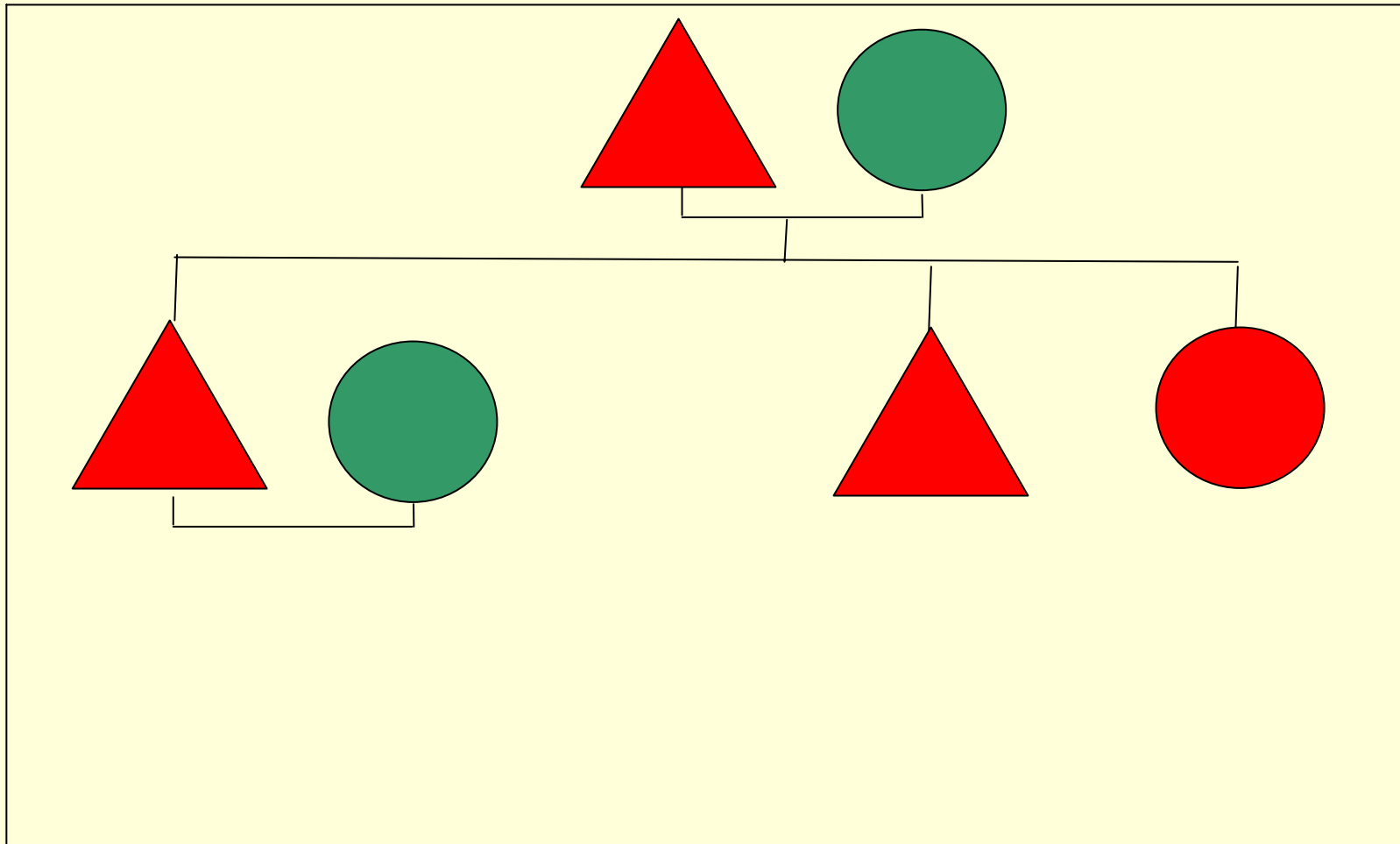
# Complex Family Structure

- For most of human history and in many parts of the world today, family systems have been far more complex than the nuclear family.
- In Croatia, an extended or joint-family household system existed with patrilocal residence of wives.
- (in english) Brothers, and sometimes cousins, lived together with their wives and unmarried sisters.

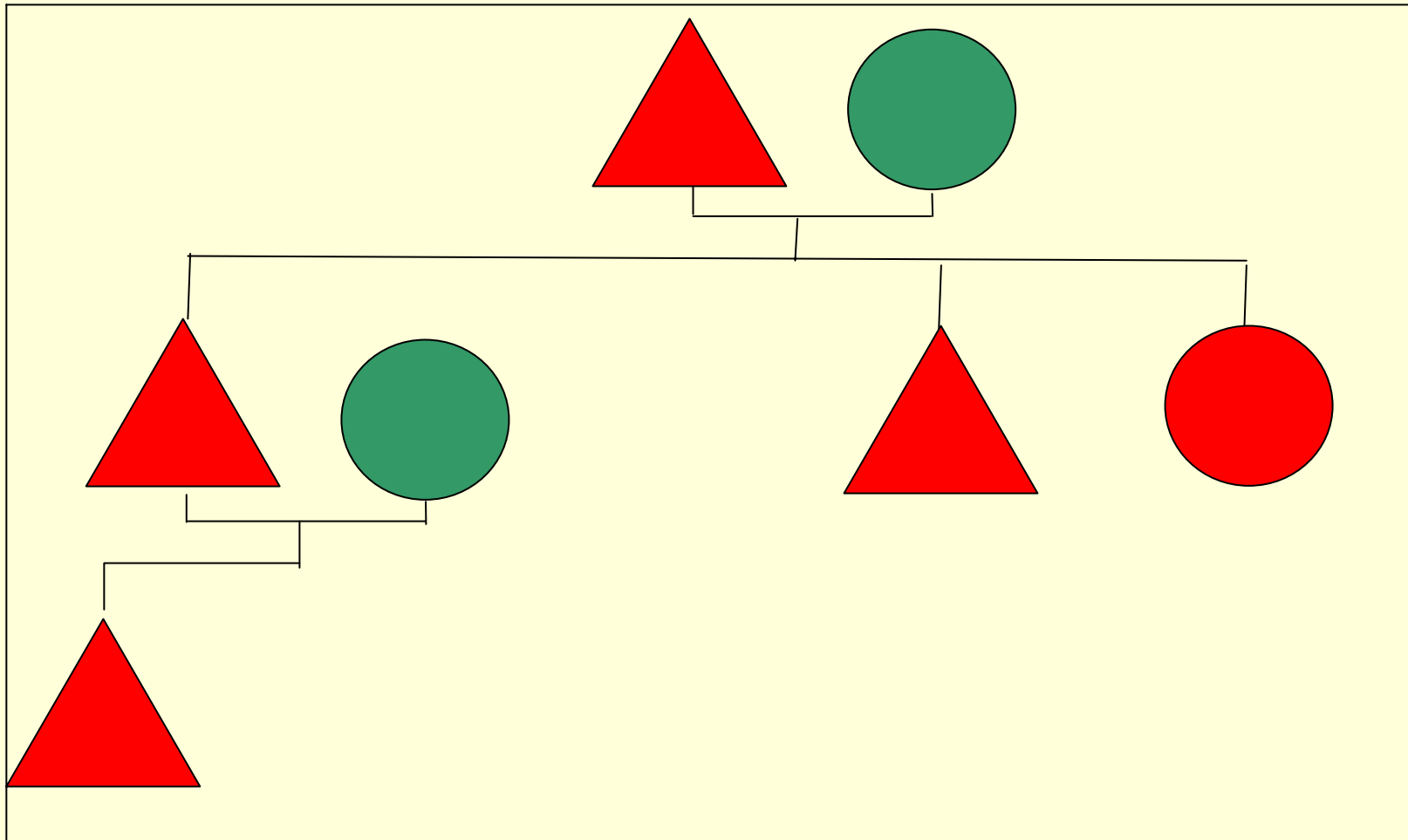
# The life-cycle of a complex household

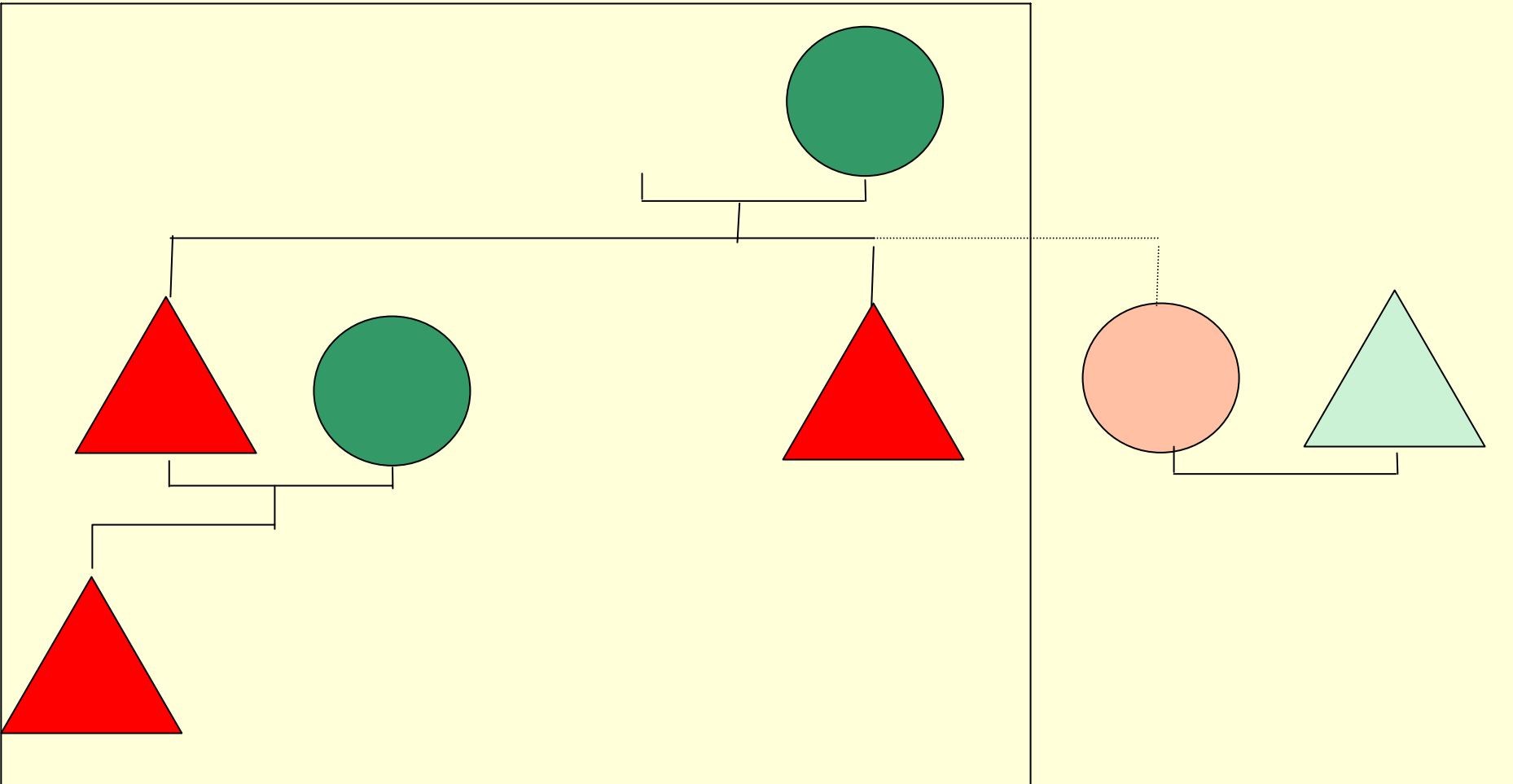


# The life-cycle of a complex household



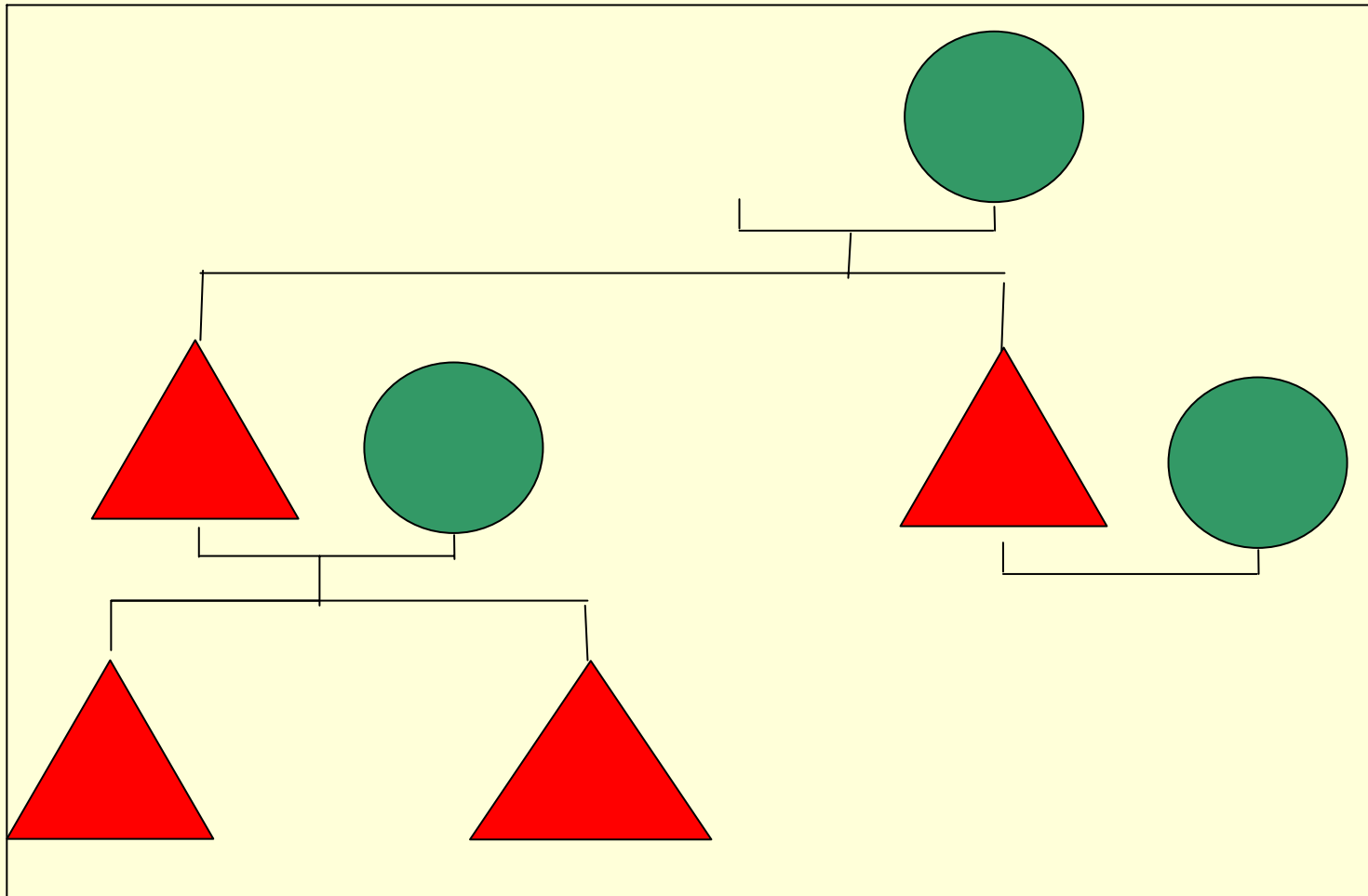
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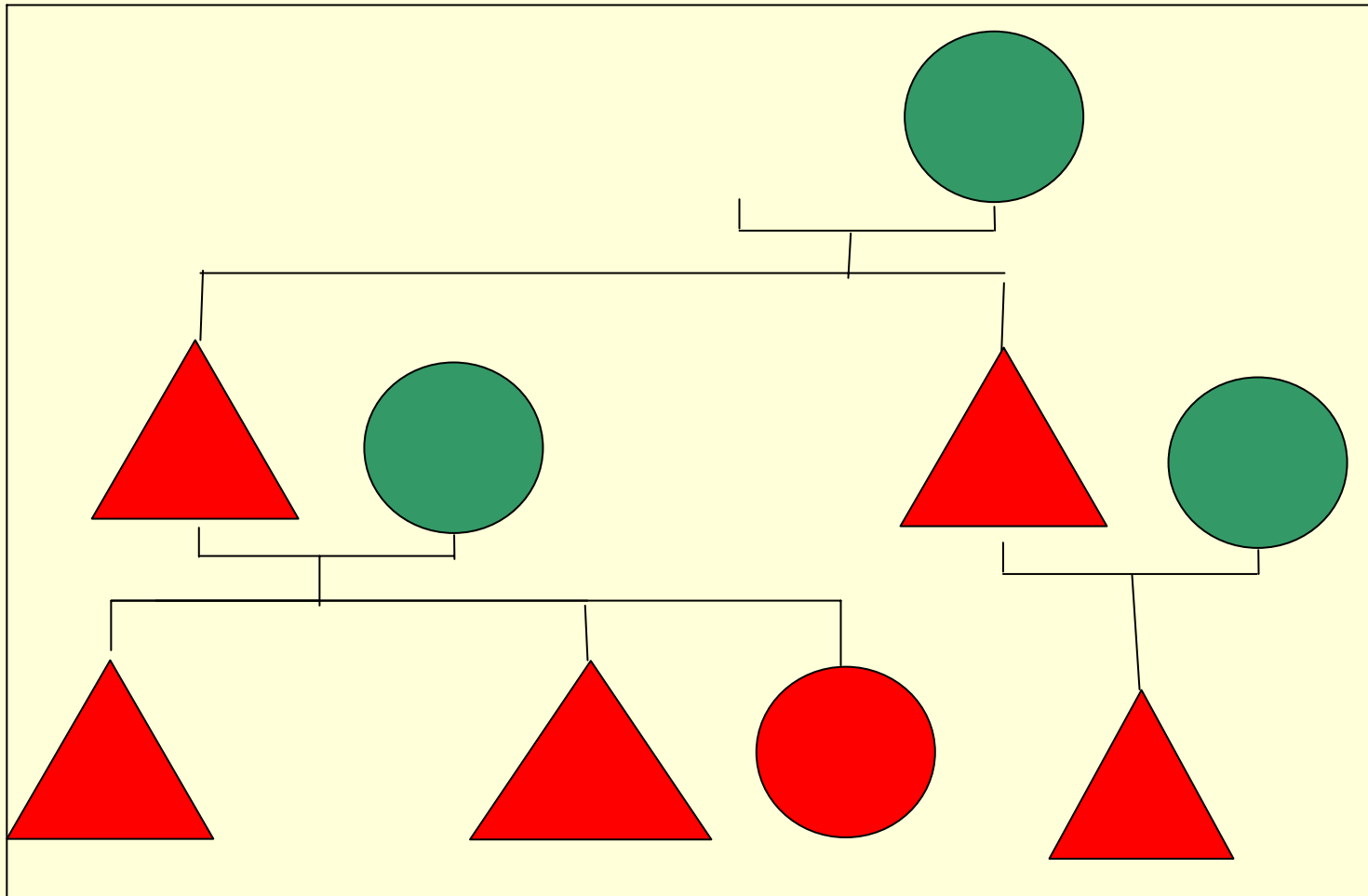
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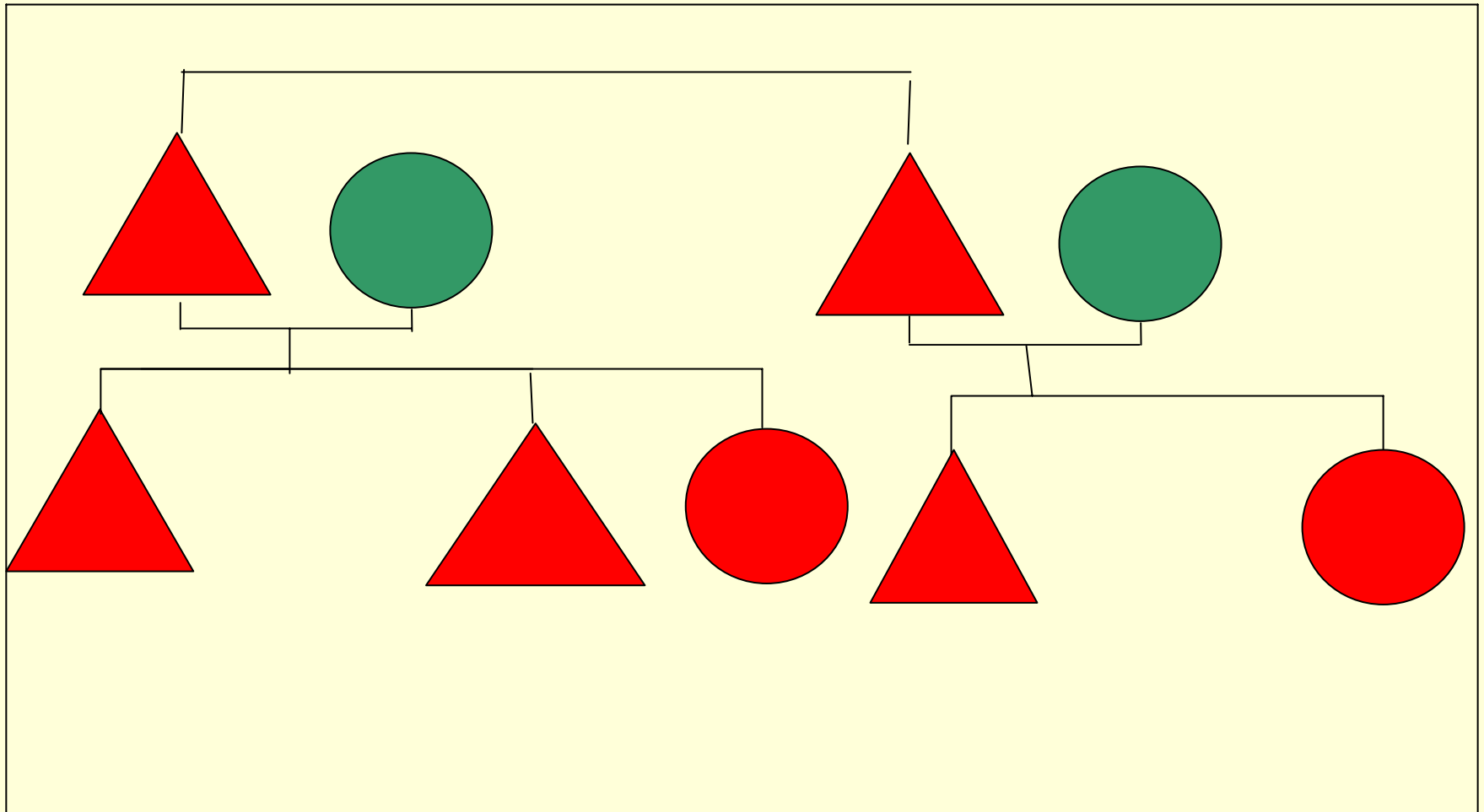
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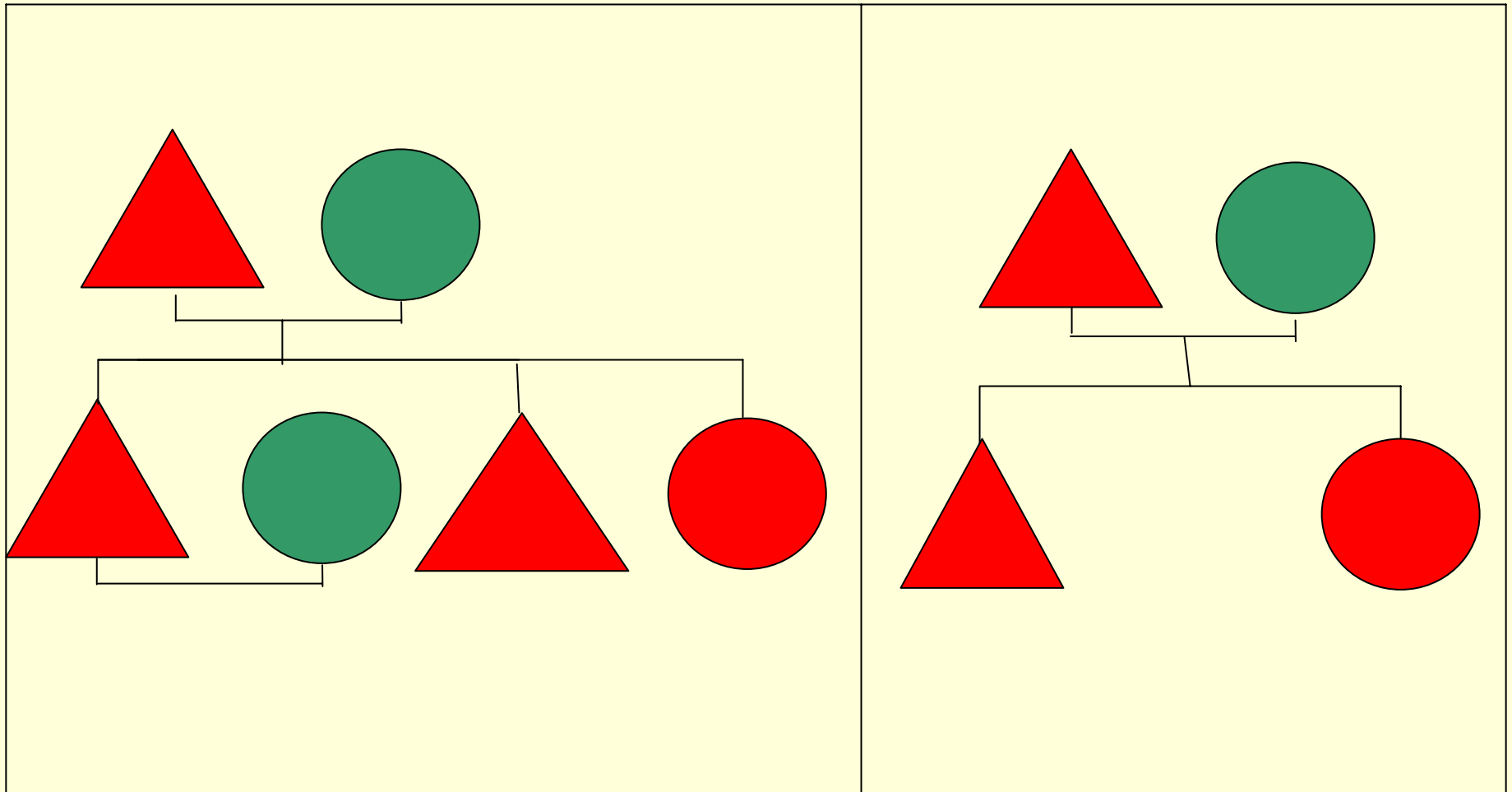
# The life-cycle of a complex household



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# The life-cycle of a complex household



# Understanding Complex Family Systems

- The life cycle of any complex household includes significant periods of nucleation.
- In pre-industrial societies, high mortality and variation in fertility will often interfere with the realization of complex family units.
- Complex, multi-generational families are the ideal, but not necessarily the most common type of family in any cross-section.

# How does this structure affect member's well being?

- Positive effect through economies of scale.
- Negative effect through the political dynamics of eventual family fission
- Stratifying effect through power differentials within the household

# Maternal Mortality

- Factors which may reduce risk of death during or after childbirth
  - The availability of care
  - Reduction in work load
- Both of these things will be mediated by a woman's husband's agnatic kinship network (and possibly by her own).

# The data

- Family reconstitution of
  - 112,181 baptisms from 1714-1898
  - 23,306 marriages from 1717-1864
  - 94,077 burials from 1717-1898
- Data are particularly good after about 1750.
- Linkage is done via a computer program which uses score optimization.
- A check with scattered village censuses shows the linkage to be quite good.



# Sample Linkage

**Birth Record**

Name: Maria Junasevic

Date: 4/9/1793

Parish: G

Village: Kovac

# Sample Linkage

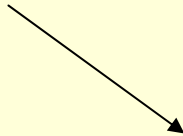
## **Birth Record**

Name: Maria Junasevic

Date: 4/9/1793

Parish: G

Village: Kovac



## **Marriage Record**

Husband: Lucas Matossevic

Wife: Maria Junasevic

Parish: G

Wife's Village: Kovac

Date: 11/29/1820

# Sample Linkage

## Birth Record

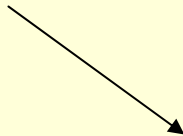
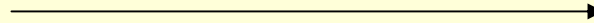
Name: Maria Junasevic  
Date: 4/9/1793  
Parish: G  
Village: Kovac

## Death Record

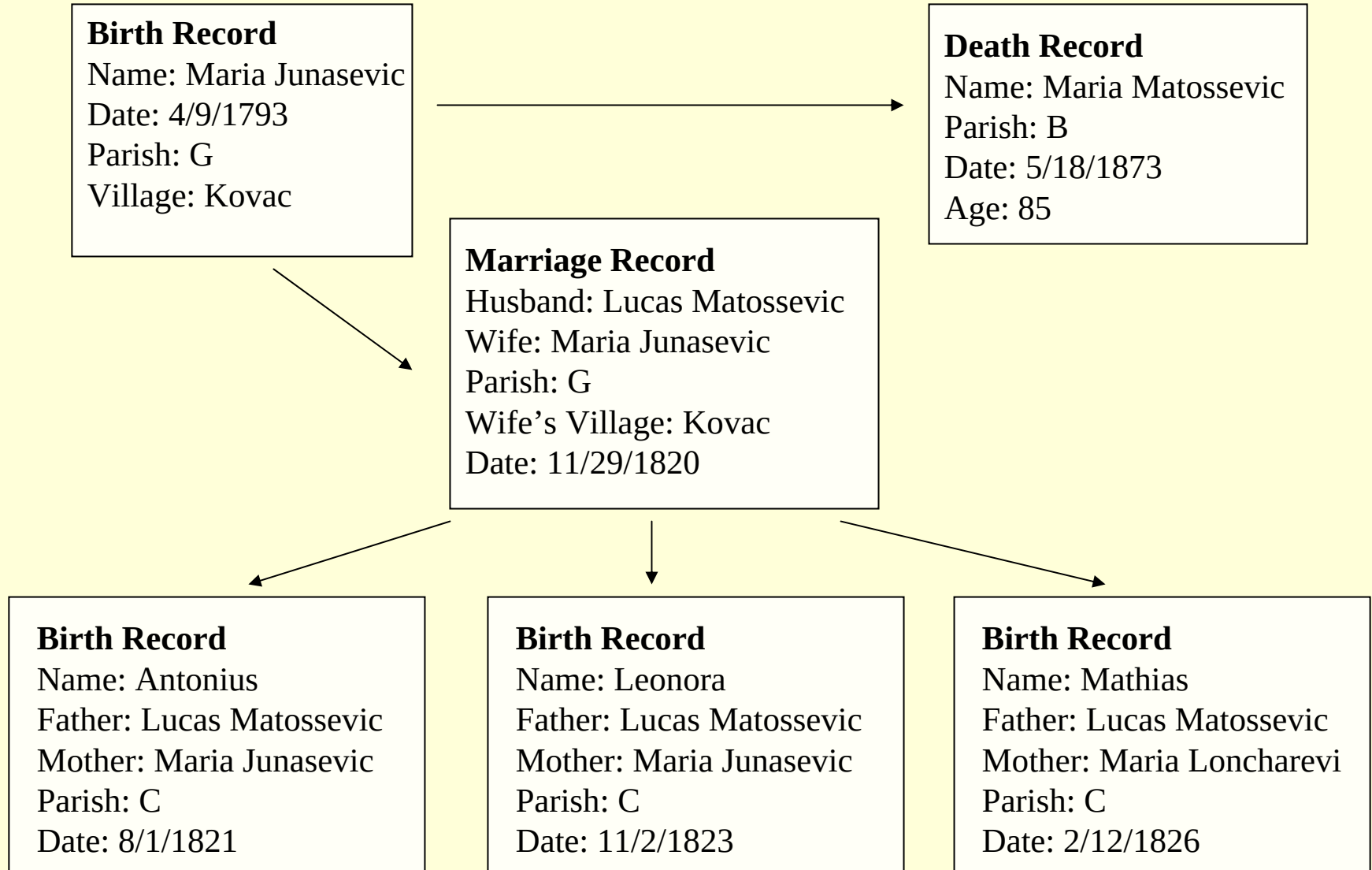
Name: Maria Matossevic  
Parish: B  
Date: 5/18/1873  
Age: 85

## Marriage Record

Husband: Lucas Matossevic  
Wife: Maria Junasevic  
Parish: G  
Wife's Village: Kovac  
Date: 11/29/1820



# Sample Linkage



# Mothers by type of linkage

| Linkage                             | Mothers | Baptisms |
|-------------------------------------|---------|----------|
| Baptism and marriage                | 8,737   | 37,757   |
| <i>Father linked to baptism</i>     | 6,261   | 27,846   |
| <i>Father not linked to baptism</i> | 2,476   | 9,911    |
| Marriage only                       | 4,465   | 18,789   |
| None                                | 16,475  | 37,712   |
| All                                 | 29,677  | 94,258   |

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# The model

- Hazard of dying within 60 days of a birth.
- Includes a random frailty term for kin network.
- For the  $j$ th birth in the  $i$ th family network:

$$\log(h(t_{ij})) = \log(w_i) + \log(\lambda_0(t_{ij})) + \beta' \mathbf{x}_{ij}$$

$\lambda_0$  is the baseline hazard,  $w_i$  is a random frailty effect, and  $\mathbf{x}_{ij}$  is a set of covariates.

# Key variables

- Crisis period
- Civil/Military distinction
- Size of affinal and consanguineal groups
- Number of husband's brother's wives
- Relative age rank in the kin network



# Male Labor Withdrawal

- Crisis periods
  - Lead to withdrawal of male labor during wartime.
  - We don't know if this particular regiment was called up for any particular crisis.
- Civil/Military distinction
  - Duties of military serfs reduced male labor inputs, even in periods of peace.

# Economies of Scale

- Number of affinal kin in the same generation:
  - Husband's brothers (hb)
  - Husband's sisters (hz)
  - Husband's brothers' wives (hbw)
- Number of consanguineal kin in the same generation:
  - Brothers (b)
  - Sisters (z)
  - Brother's wives (bw)

# Political Division

- The number of hbw's in the household relative to other types of kin will adversely affect health, because hbw's will be more likely to withhold labor and care.
  - The number of husband's brothers' wives contributes both to economies of scale and division.

# Stratification

- The age rank of a woman's brother in the household will affect the availability of care and resources.
  - $W$  is the number of wives in the household
  - $A$  is the absolute age rank (1 is the most senior)
  - $R$  is relative age rank:
$$R = (W - A) / (W - 1)$$
  - 1 is senior wife, 0 is most junior. Intermediate wives are scaled between 1 and 0.
  - Women without hbw are assigned a rank of 1, and given a dummy variable.

# Other Controls

- Physiological variables
  - Age
  - Parity
  - Previous and mean birth interval
  - Twinning
  - Background mortality
- Temporal variables
  - Linear time trend
  - Parity 1 \* linear time trend interaction

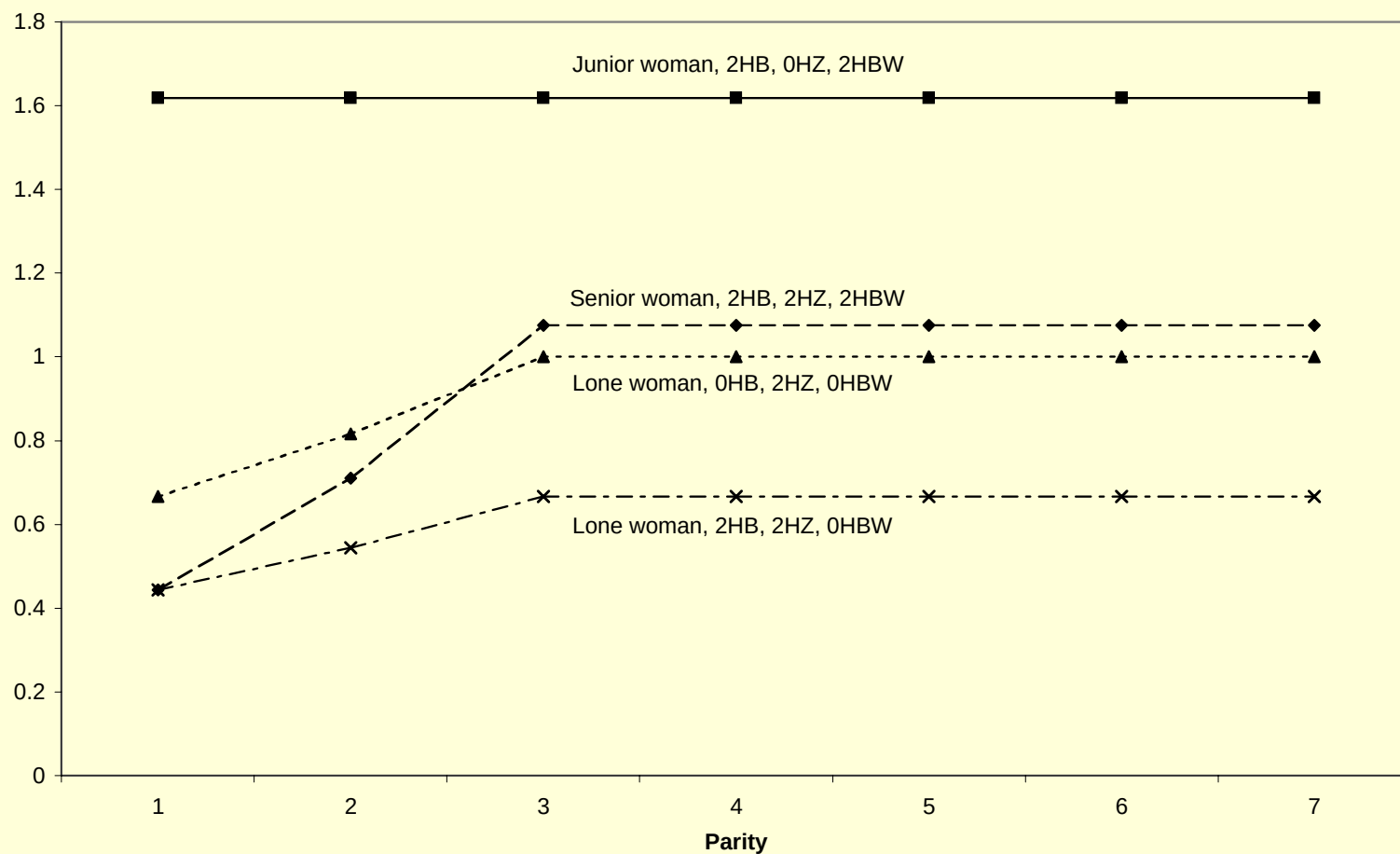
# Maternal Mortality Results

| Variable           | Coefficient | SE       |
|--------------------|-------------|----------|
| Crisis Period      | 0.313       | 0.105*** |
| Parish             |             |          |
| Civil Parish (ref) | -           |          |
| B                  | 0.546       | 0.313*   |
| G                  | 0.109       | 0.226    |
| O                  | 0.339       | 0.250‡   |
| S                  | 0.353       | 0.207*   |
| P                  | 0.346       | 0.267    |
| V                  | -0.054      | 0.255    |

# Maternal Mortality Results (cont)

| Variable             | Coefficient | SE      |
|----------------------|-------------|---------|
| # of Adult Affines   | -0.238      | 0.083** |
| # Adult Consanguines | -0.032      | 0.038   |
| # of HBW             | 0.477       | 0.236*  |
| Relative Rank        | -0.396      | 0.249‡  |
| No HBW               | -0.047      | 0.338   |
| Variance of frailty  | 0.489       |         |

# Change in relative risk over family life-cycle





# Infant and Child Mortality

- Additional issue of the age and gender composition of existing siblings and cousins.
  - Positive effect through economies of scale with younger children, and childrearing assistance from older children.
  - Negative effect through resource shortages, particularly as cousins age and household fission approaches (inheritance shares).

# The model

- Piecewise Constant Exponential Model
  - 0-1 months, 1-6 months, 6-12 months, 12-24 months, 24-60 months

$$\log(h_{it}) = \log(\lambda_t) + \beta' \mathbf{x}_{it}$$

- We want to add two, and possibly three, kinds of frailty
  - Mother effect
  - **Kin network effect**
  - Village effect

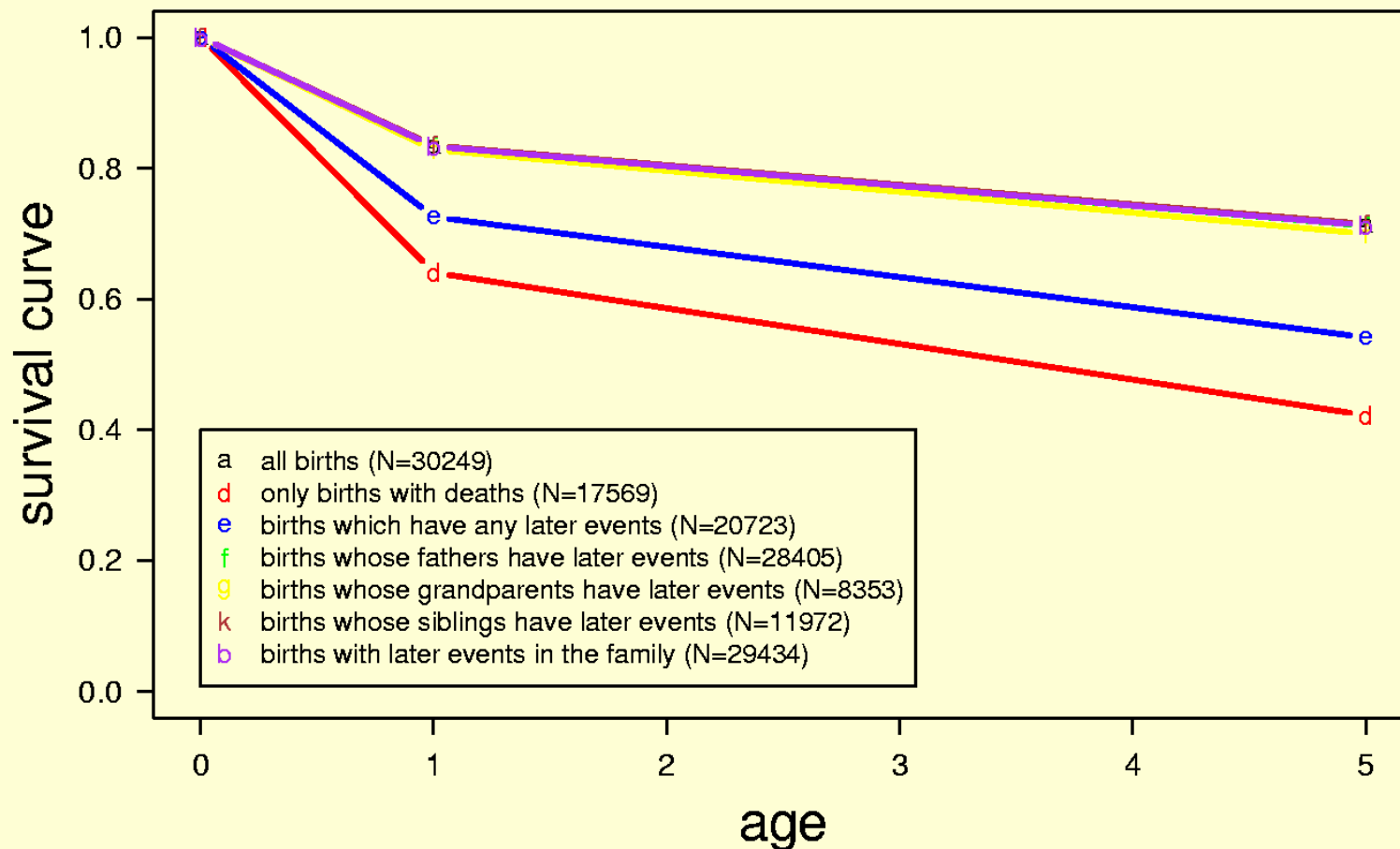
# Key Variables

- Number of siblings and cousins in age categories:
  - 0-5 years
  - 5-10 years
  - 10 years until marriage or censoring at 25.

# Censoring Observations

- In cases where no death record is recorded, how do you censor an observation?
- We assume all children without an observed death survive to five years of age.
- This approach biases overall death rates, but is hopefully less problematic for covariates.

# Survival curves under different censoring rules



# Preliminary Results

| Variable           | Coefficient | SE       |
|--------------------|-------------|----------|
| Crisis Period      | 0.190       | 0.021*** |
| Parish             |             |          |
| Civil Parish (ref) | -           |          |
| B                  | 0.513       | 0.053*** |
| G                  | 0.373       | 0.034*** |
| O                  | 0.200       | 0.044*** |
| S                  | 0.564       | 0.042*** |
| P                  | 0.555       | 0.033*** |
| V                  | 0.393       | 0.036*** |

# Preliminary Results (cont)

| Variable             | Coefficient | SE    |
|----------------------|-------------|-------|
| # of Adult Affines   | -0.005      | 0.011 |
| # Adult Consanguines | 0.003       | 0.006 |
| # of HBW             | -0.020      | 0.040 |
| Relative Rank        | 0.004       | 0.037 |
| No HBW               | -0.035      | 0.050 |
| Variance of frailty  | 0.082       |       |

# Preliminary Results (cont)

| Variable       | Coefficient | SE       |
|----------------|-------------|----------|
| # of Siblings  |             |          |
| 0-5 years old  | -0.164      | 0.021*** |
| 5-10 years old | -0.013      | 0.020    |
| 10+ years old  | -0.063      | 0.020**  |
| # of Cousins   |             |          |
| 0-5 years old  | 0.028       | 0.017    |
| 5-10 years old | 0.076       | 0.018*   |
| 10+ years old  | -0.017      | 0.011    |



# Conclusions

- Family structure has complex, multi-dimensional relationship with members' well-being
- Family planning policy would benefit from an understanding of this relationship in many developing nations.